

*Discussion of:*  
Asset Prices with Fading Memory  
Stefan Nagel and Zhengyang Xu

Kent Daniel<sup>†</sup>

<sup>†</sup>Columbia Business School & NBER

2018 Fordham Rising Stars Conference  
May 11, 2018



# Introduction

- The paper has two distinct parts:
  - a simple model with some empirical work
  - a more sophisticated Bayesian learning model.
- I'm going to concentrate on the first part.
- In both models agents learn the underlying cashflow growth rate by observing realized cashflow growth rates.
  - As a result of experiential learning, investors overreact to recent growth rates.
  - Agents get no other information.
- I'm going to argue that you need other shocks to explain market returns.

# Basic Model - Cash flow process

- The cashflow  $C_t$  from the endowment (market) follows GBM with constant drift  $\mu$ :

$$\Delta c_t = \mu + \epsilon_t$$

where  $c_t = \log(C_t)$  and  $\epsilon_t \sim iid \mathcal{N}(0, \sigma_c^2)$

- However, based on Malmendier and Nagel (2016), the *average* agent's belief about  $\mu$ ,  $\tilde{\mu}$ , follows:

$$\tilde{\mu}_{t+1} = \tilde{\mu}_t + \nu (\Delta c_{t+1} - \tilde{\mu}_t)$$

- MN (2016) estimate  $\nu = 0.018$ /quarter for inflation data.
- implying that:

$$\tilde{\mu}_t = \sum_{j=0}^{\infty} \underbrace{\nu(1-\nu)^j}_{w_j} \Delta c_{t-j}$$

# Basic Model - Cash flow process

- The cashflow  $C_t$  from the endowment (market) follows GBM with constant drift  $\mu$ :

$$\Delta c_t = \mu + \epsilon_t$$

where  $c_t = \log(C_t)$  and  $\epsilon_t \sim iid \mathcal{N}(0, \sigma_c^2)$

- However, based on Malmendier and Nagel (2016), the *average* agent's belief about  $\mu$ ,  $\tilde{\mu}$ , follows:

$$\tilde{\mu}_{t+1} = \tilde{\mu}_t + \nu (\Delta c_{t+1} - \tilde{\mu}_t)$$

- MN (2016) estimate  $\nu = 0.018$ /quarter for inflation data.
- implying that:

$$\tilde{\mu}_t = \sum_{j=0}^{\infty} \underbrace{\nu(1-\nu)^j}_{w_j} \Delta c_{t-j}$$

# Basic Model - Cash flow process

- The cashflow  $C_t$  from the endowment (market) follows GBM with constant drift  $\mu$ :

$$\Delta c_t = \mu + \epsilon_t$$

where  $c_t = \log(C_t)$  and  $\epsilon_t \sim iid \mathcal{N}(0, \sigma_c^2)$

- However, based on Malmendier and Nagel (2016), the *average* agent's belief about  $\mu$ ,  $\tilde{\mu}$ , follows:

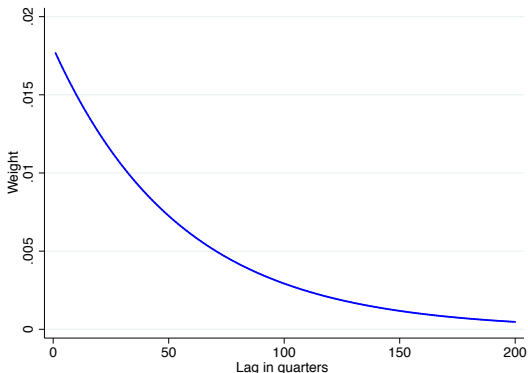
$$\tilde{\mu}_{t+1} = \tilde{\mu}_t + \nu (\Delta c_{t+1} - \tilde{\mu}_t)$$

- MN (2016) estimate  $\nu = 0.018$ /quarter for inflation data.
- implying that:

$$\tilde{\mu}_t = \sum_{j=0}^{\infty} \underbrace{\nu(1-\nu)^j}_{w_j} \Delta c_{t-j}$$

# Weighting Function, $\nu = 0.018$

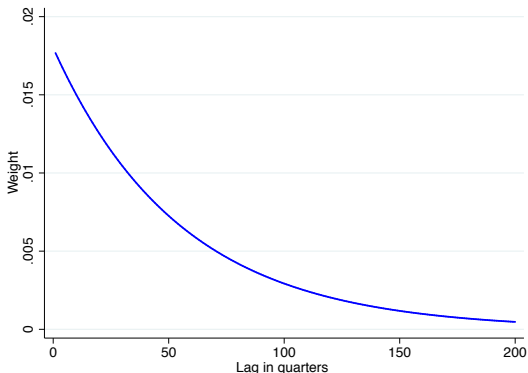
- That is,  $\tilde{\mu}_t = \sum_{j=0}^{\infty} w_j \Delta c_{t-j}$ , where  $w_j$  looks like:



- half life is  $\frac{\log(0.5)}{\log(1-\nu)} = 38.2$  quarters ( $\sim 10$  years)
  - $\nu \rightarrow 0 \Rightarrow$  “rationality” (*i.e.*, no “fading”)

# Weighting Function, $\nu = 0.018$

- That is,  $\tilde{\mu}_t = \sum_{j=0}^{\infty} w_j \Delta c_{t-j}$ , where  $w_j$  looks like:



- half life is  $\frac{\log(0.5)}{\log(1-\nu)} = 38.2$  quarters ( $\sim 10$  years)
  - $\nu \rightarrow 0 \Rightarrow$  “rationality” (*i.e.*, no “fading”)

# Basic Model - Pricing

- Representative agent in model sets price to equal PV of future CFs, using constant discount rate of

$$\tilde{\mathbb{E}}_t r_{t+1} = \theta + r_f$$

- However, agent (mistakenly) extrapolates recent cashflow growth to infer  $\mu$ .
- Using a Campbell and Shiller (1988) log-linearization:

$$r_{t+1} - \tilde{\mathbb{E}}_t r_{t+1} = \left(1 + \left(\frac{\rho}{1-\rho}\right)\nu\right) (\Delta c_{t+1} - \tilde{\mu}_t)$$

and

$$\mathbb{E}_t r_{t+1} - \underbrace{\tilde{\mathbb{E}}_t r_{t+1}}_{r_f + \theta} = \left(1 + \left(\frac{\rho}{1-\rho}\right)\nu\right) (\mu - \tilde{\mu}_t)$$

implying a negative relationship between recent cashflow growth and future abnormal returns.



# Basic Model - Pricing

- Representative agent in model sets price to equal PV of future CFs, using constant discount rate of

$$\tilde{\mathbb{E}}_t r_{t+1} = \theta + r_f$$

- However, agent (mistakenly) extrapolates recent cashflow growth to infer  $\mu$ .
- Using a Campbell and Shiller (1988) log-linearization:

$$r_{t+1} - \tilde{\mathbb{E}}_t r_{t+1} = \left(1 + \left(\frac{\rho}{1-\rho}\right)\nu\right) (\Delta c_{t+1} - \tilde{\mu}_t)$$

and

$$\mathbb{E}_t r_{t+1} - \underbrace{\tilde{\mathbb{E}}_t r_{t+1}}_{r_f + \theta} = \left(1 + \left(\frac{\rho}{1-\rho}\right)\nu\right) (\mu - \tilde{\mu}_t)$$

implying a negative relationship between recent cashflow growth and future abnormal returns.

# Basic Model - Pricing

- Representative agent in model sets price to equal PV of future CFs, using constant discount rate of

$$\tilde{\mathbb{E}}_t r_{t+1} = \theta + r_f$$

- However, agent (mistakenly) extrapolates recent cashflow growth to infer  $\mu$ .
- Using a Campbell and Shiller (1988) log-linearization:

$$r_{t+1} - \tilde{\mathbb{E}}_t r_{t+1} = \left(1 + \left(\frac{\rho}{1-\rho}\right) \nu\right) (\Delta c_{t+1} - \tilde{\mu}_t)$$

and

$$\mathbb{E}_t r_{t+1} - \underbrace{\tilde{\mathbb{E}}_t r_{t+1}}_{r_f + \theta} = \left(1 + \left(\frac{\rho}{1-\rho}\right) \nu\right) (\mu - \tilde{\mu}_t)$$

implying a negative relationship between recent cashflow growth and future abnormal returns.

# Estimating $\tilde{\mu}_t$

- The authors don't use cashflows to estimate  $\tilde{\mu}_t$ .
- They instead use historical returns on the market. Effectively,  
$$\tilde{\mu}_{r,t} = \sum_{j=0}^{\infty} w_j r_{t-j}$$
- Reasons:
  - ① To start in 1926, would need consumption going back to 1876.
  - ② "...dividends are influenced by shifts in payout policy that can distort estimates of  $\tilde{\mu}$  constructed from dividend growth rates."
  - ③ The authors simulate  $\tilde{\mu}$  and  $\tilde{\mu}_r$  (under the null) and show that they are highly correlated.
- A concern is that price shocks will reflect *all* information prices/discount rates.
  - How can we confirm the information that is causing  $\mathbb{E}[r]$ s to change is cashflow innovations?

# Estimating $\tilde{\mu}_t$

- The authors don't use cashflows to estimate  $\tilde{\mu}_t$ .
- They instead use historical returns on the market. Effectively,  
$$\tilde{\mu}_{r,t} = \sum_{j=0}^{\infty} w_j r_{t-j}$$
- Reasons:
  - ① To start in 1926, would need consumption going back to 1876.
  - ② "...dividends are influenced by shifts in payout policy that can distort estimates of  $\tilde{\mu}$  constructed from dividend growth rates."
  - ③ The authors simulate  $\tilde{\mu}$  and  $\tilde{\mu}_r$  (under the null) and show that they are highly correlated.
- A concern is that price shocks will reflect *all* information prices/discount rates.
  - How can we confirm the information that is causing  $\mathbb{E}[r]$ s to change is cashflow innovations?

# DP decomposition

- I'll show a set of regressions. Data is from Shiller, over the 1946-2014 sample.
- The dependent variable is always the annual real returns on the S&P 500 ( $R_{t+1}$ )
- The forecasting variables I'll use are:
  - ①  $dp$ : log of preceding year's dividend ( $D_t$ ), scaled by this year's price ( $P_t$ )
  - ②  $dpL$ :  $dp$ , lagged 10 years.
  - ③  $\Delta d$ : change in the log dividend over the last 10 years.
  - ④  $\Delta p$ : change in the log price over the last 10 years.
  - ⑤  $S$ : Baker and Wurgler (2000) equity share

# An Information Decomposition

## OLS Regression Results

```

=====
Dep. Variable:                R      R-squared:                0.066
Model:                        OLS    Adj. R-squared:          0.052
No. Observations:            67     AIC:                    -51.62
Df Residuals:                 65     BIC:                    -47.21
Df Model:                      1
Covariance Type:              HAC
=====

```

```

=====
              coef      std err          z      P>|z|      [0.025      0.975]
-----
const          0.4165      0.157      2.657      0.008      0.109      0.724
dp              0.0983      0.046      2.128      0.033      0.008      0.189
=====

```

## An Information Decomposition

## OLS Regression Results

```

=====
Dep. Variable:                R      R-squared:                0.026
Model:                        OLS    Adj. R-squared:          0.011
Method:                       Least Squares  F-statistic:            2.210
Date:                          Thu, 10 May 2018  Prob (F-statistic):     0.142
Time:                          09:57:50    Log-Likelihood:        26.393
No. Observations:             67      AIC:                    -48.79
Df Residuals:                 65      BIC:                    -44.38
Df Model:                      1
Covariance Type:              HAC
=====

```

	coef	std err	z	P> z	[0.025	0.975]
const	0.0588	0.022	2.676	0.007	0.016	0.102
Delta-d	0.1254	0.084	1.487	0.137	-0.040	0.291

- the point estimate on the  $\Delta d$  coefficient is positive, not negative.
  - However, it is statistically insignificant.

# dp decomposition

- Consider the identity (like that in Daniel and Titman (2006)):

$$dp_t = dp_{t-10} + \Delta d_{t-10,t} - \Delta p_{t-10,t}$$

- In words, if the market has a high  $dp$  today, there are three possibilities:
  - ④ It was high  $dp$  10 years ago.
  - ②  $\Delta d$  was positive.
  - ③  $\Delta p$  was negative.
- At least post-WWII,  $dp$  forecasts the market.
  - *Which of the three components forecasts the market?*



# An Information Decomposition

## OLS Regression Results

```

=====
Dep. Variable:                R    R-squared:                0.111
Model:                        OLS  Adj. R-squared:           0.069
Method:                        Least Squares  F-statistic:              2.882
Date:                          Thu, 10 May 2018  Prob (F-statistic):      0.0427
Time:                          09:57:50  Log-Likelihood:          29.465
No. Observations:              67    AIC:                     -50.93
Df Residuals:                  63    BIC:                     -42.11
Df Model:                       3
Covariance Type:              HAC
=====

```

```

=====
              coef      std err          z      P>|z|      [0.025      0.975]
-----+-----
const          0.4635      0.165      2.814      0.005      0.141      0.786
dpL            0.1192      0.051      2.353      0.019      0.020      0.219
Delta-d        0.2698      0.117      2.300      0.021      0.040      0.500
Delta-p       -0.1077      0.051     -2.123      0.034     -0.207     -0.008
=====

```

- Note that the coefficient on  $\Delta d$  is again **positive**, and now statistically significant.
  - Suggests that  $\Delta d$  is not just “noise” w.r.t returns.

## dp decomposition

- We can also break the market “return” into the part explained by cashflow changes, and the component that isn’t ( $\epsilon$ ).

$$\Delta p_{t-10,t} = a \cdot dp_{t-10} + b \cdot \Delta d_{t-10,t} + \epsilon_{t-10,t}$$

- $\epsilon_{t-10,t}$  is the price change over the last 10 years that can’t be explained by the growth rate of dividends.
  - The regression  $R_{adj}^2 = 51.4\%$
  - $t(b = 0) = 6.7$ .

## An Information Decomposition

## OLS Regression Results

```

=====
Dep. Variable:          Delta-p      R-squared:                0.529
Model:                  OLS          Adj. R-squared:           0.514
Method:                 Least Squares  F-statistic:              30.51
Date:                   Thu, 10 May 2018  Prob (F-statistic):       4.94e-10
Time:                   10:45:20      Log-Likelihood:           -29.072
No. Observations:      67           AIC:                      64.14
Df Residuals:          64           BIC:                      70.76
Df Model:               2
Covariance Type:      HAC
=====

```

	coef	std err	z	P> z	[0.025	0.975]
const	1.7185	0.357	4.808	0.000	1.018	2.419
dpL	0.4961	0.109	4.537	0.000	0.282	0.710
Delta-d	1.4854	0.221	6.734	0.000	1.053	1.918

```

=====

```

- $R_{adj.}^2 = 51.4\% \Rightarrow \rho \approx 0.7,$

## An Information Decomposition

## OLS Regression Results

```

=====
Dep. Variable:          R      R-squared:          0.111
Model:                 OLS    Adj. R-squared:     0.069
Method:                Least Squares    F-statistic:        2.882
Date:                  Thu, 10 May 2018    Prob (F-statistic): 0.0427
Time:                  11:21:31    Log-Likelihood:     29.465
No. Observations:      67      AIC:                -50.93
Df Residuals:          63      BIC:                -42.11
Df Model:              3
Covariance Type:      HAC
=====

```

```

=====
              coef      std err          z      P>|z|      [0.025      0.975]
-----
const          0.2784      0.137          2.030      0.042      0.010      0.547
dpL            0.0658      0.043          1.544      0.123     -0.018      0.149
Delta-d        0.1098      0.083          1.328      0.184     -0.052      0.272
resid         -0.1077      0.051         -2.123      0.034     -0.207     -0.008
=====

```

- The coefficient on *resid* is exactly the same as in the previous regression.
- The coefficients on  $dp_{t-10}$  and  $\Delta d$  are what they would be were *resid* not included in the regression.

# An Information Decomposition

## OLS Regression Results

```

=====
Dep. Variable:          R      R-squared:          0.166
Model:                  OLS    Adj. R-squared:    0.104
Method:                 Least Squares    F-statistic:       2.878
Date:                   Thu, 10 May 2018    Prob (F-statistic): 0.0311
Time:                   09:57:51    Log-Likelihood:    28.634
No. Observations:      59    AIC:               -47.27
Df Residuals:          54    BIC:               -36.88
Df Model:               4
Covariance Type:      HAC
=====

```

```

=====
              coef      std err          z      P>|z|      [0.025      0.975]
-----
const          0.6982      0.227       3.077     0.002     0.253     1.143
dpL            0.1585      0.068       2.318     0.020     0.024     0.293
Delta-d        0.3394      0.124       2.728     0.006     0.096     0.583
Delta-p       -0.1589      0.057      -2.765     0.006    -0.272    -0.046
S              -0.5252      0.283      -1.858     0.063    -1.079     0.029
=====

```

# References I

- Baker, Malcolm, and Jeffrey Wurgler, 2000, The equity share in new issues and aggregate stock returns, *Journal of Finance* 55, 2219–2257.
- Campbell, John Y., and Robert J. Shiller, 1988, The dividend-price ratio and expectations of future dividends and discount factors, *Review of Financial Studies* 1, 195–228.
- Daniel, Kent D., and Sheridan Titman, 2006, Market reactions to tangible and intangible information, *Journal of Finance* 61, 1605–1643.